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Statistical design and analysis of field trials on
non-target species
Alternatives to OECD-decision tree

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Structure I

- Possible effect of new agro-chemicals on non-target species in randomized field trials ⇒ small sample sizes issues inherently (not only power!)
- Design: i) CRD, ii) $n_i = 6$ pre-defined by OECD guidance (follow strictly!), iii) C+, C_1 , ..., C_3 , C+, iv) multiple sampling times T_0 , T_t , v) many primary multiple count endpoints (mostly solicited, unsolicited) and their taxonomic aggregations (serious multiplicity issue!)
- Actually a proof of safety, i.e. non-inferiority tests (clearly a directional hypothesis) or better confidence limits. But species-specific tolerable thresholds ξ_p unknown. Therefore proof of hazard with all the nice confusions, e.g. *'The absence of evidence is no evidence of absence'* (Altman/Bland)

An example I

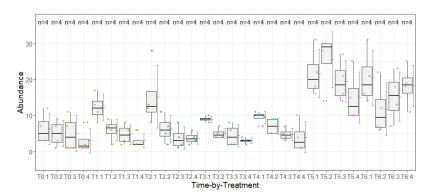
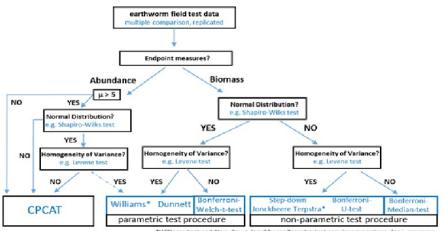


Figure: Ascidea abundance example data

OECD decision tree I

OECD decision tree II



*Williams test and Step-down-Jonckheere Terpstra test requires monotone dose-response

Issues of the OECD decision tree I

- 2 primary endpoints- differently scaled
- ullet Abundances: cutpoint $\mu=5$ taken from a very old textbook. Hard to defined so simple
- General pre-test/post test dilemma: lack-of-fit tests controlling the less relevant error rate. Equivalence tests needed with a pre-defined tolerance threshold...
- Both test on normality and variance homogeneity too low power for small n_i's
- Alternative test not comparable (e.g. quite different effect sizes)
- Such decision trees should not be recommended at all.
 Alternative: well-chosen, robust tests

Why several concentrations? I

- Assuming Paracelsus law: All things are poison and nothing is without poison. Solely the dose determines that a thing is not a poison
- OECD propose statistically two quite different approaches: NOEC, BMD
- See Zhenglei's talk on BMD in ecotox
- Both approaches with pros and cons. In the following NOEC only here

Estimate NOEC I

- NOEC depends on the unadjusted effect size δ . Yes, it should. But also on $n_i, s_i, C_k, \Delta C_i, \ldots$
- OECD design recommendation represents a sort of standardization of this point-zero-null-hypotheses tests
- First impulse: ordered concentrations require order restricted tests (to increase power)
- BUT-they use aggregations of C_i and this biased NOEC estimation. Williams trend test as a simple example for a specific decreasing plateau-shaped profile:

$$H_1: \mu_0 = 5 > \mu_1 = 3 < \mu_2 = 4 > \mu_3 = 3.5$$
:

| $\pi_{Du_1=0}$ | $\pi_{Du_{2-0}}$ | $\pi_{Du_{3-0}}$ | $\pi_{Wi_{1-0}}$ | $\pi_{Wi_{2-0}}$ | $\pi_{W_{i_3-0}}$ |
|----------------|------------------|------------------|------------------|------------------|-------------------|
| 0.90 | 0.11 | 0.66 | 0.71 | 0.71 | 0.83 |

Table: Per-pair powers of Dunnett and Williams test for a simulated specific plateau-shaped alternative

Estimate NOEC II

- Notice, the per-pair power is relevant for NOEC, whereas any-pairs power is commonly published
- Alternative: Dunnett test.
- But Dunnett test may be biased when heterogeneous variances occur- problematic when in the non-NOEC concentrations
- An example- compared with unbiased Welch-type-df modification:

Estimate NOEC III

| NOEC | s _i ↑ | Original Dunnett | | | Welch-type Dunnett | | |
|------|------------------|------------------|-----------|-----------|--------------------|-----------|-----------|
| | | D_{1-0} | D_{2-0} | D_{3-0} | W_{1-0} | W_{2-0} | W_{3-0} |
| 2 | - | 0.02 | 0.02 | 0.89 | 0.02 | 0.02 | 0.85 |
| 2 | 3 | 0.00 | 0.00 | 0.35 | 0.02 | 0.02 | 0.15 |
| 2 | 2 | 0.00 | 0.07 | 0.23 | 0.02 | 0.02 | 0.83 |
| 2 | 1 | 0.07 | 0.01 | 0.24 | 0.02 | 0.01 | 0.82 |
| 2 | 0 | 0.07 | 0.07 | 0.37 | 0.03 | 0.03 | 0.21 |
| 1 | - | 0.02 | 0.89 | 0.88 | 0.93 | 0.84 | 0.83 |
| 1 | 3 | 0.00 | 0.24 | 0.38 | 0.02 | 0.83 | 0.14 |
| 1 | 2 | 0.00 | 0.36 | 0.22 | 0.02 | 0.13 | 0.83 |
| 1 | 1 | 0.08 | 0.24 | 0.24 | 0.02 | 0.82 | 0.82 |
| 1 | 0 | 0.07 | 0.36 | 0.37 | 0.03 | 0.19 | 0.20 |
| 0 | - | 0.90 | 0.91 | 0.89 | 0.97 | 0.85 | 0.85 |
| 0 | 3 | 0.26 | 0.24 | 0.36 | 0.82 | 0.84 | 0.15 |
| 0 | 2 | 0.23 | 0.36 | 0.23 | 0.82 | 0.14 | 0.81 |
| 0 | 1 | 0.38 | 0.25 | 0.24 | 0.15 | 0.82 | 0.82 |
| 0 | 0 | 0.41 | 0.36 | 0.35 | 0.20 | 0.19 | 0.20 |

Table: Simulated per-pair power estimates of Dunnett and Welch-type Dunnett procedure for selected NOEC's and patterns of variance heterogeneity: fair power loss; bias

Estimate NOEC IV

 Recommendation: Estimate NOEC neither by Williams, nor by original Dunnett

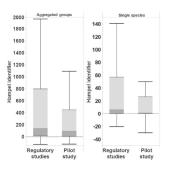
use Welch/sandwich modifications!

Count data issues I

- Two primary endpoints per taxonomic level: i) abundance, ii) biomass. 1st count, 2nd continuous
- Empirically heterogeneous variances are observed
- For count data we have TWO related effects:
 - overdispersion for the count variable in itself
 - varying overdispersion with concentration -analogous to heteroscedasticity

Count data issues II

- UBA itself:



- Really challenging: evaluation of overdispersed count data in low n_i and k+1 designs!

OECD proposed test: CPCAT I

- CPCAT based on 2 principles:
 - **1** CP: Closed testing procedure for comparing C_i 's vs. C-
 - ② CAT: permutative version of LR-test for comparing Poisson distributed counts
- CPCAT's main idea: feasible for small sample sizes count data
- CP-part ok, but does not provide (interpretable) confidence intervals
- CAT part problematic when data overdispersed.... and some data are severe overdispersed
- OECD: 'The theoretical distribution assumption of earthworm abundance field test data follows a Poisson model'. Violates basic stats paradigm. 'All models are wrong, some are useful' G.Box

OECD proposed test: CPCAT II

- Aggregated data: i) over 4 traps, ii) over taxa.
 - ➤ **Stats**: The sum over Poisson variables is only Poisson for complete independence [1]. But they are dependent per definition
 - ► **Empirical**: UBA data reveal both under and overdispersion, rarely near-to-Poisson data in histor. data [2]
 - ► CP-CAT: 'over-dispersion reduced the statistical power of the CPCAT' Lehmann et al. 2018.

Special features:

- ▶ small n_i
- not just overdispersion, but concentration-specific dispersions similar to variance heterogeneity in Gaussian models

| Distrib. | MLT | Nonpar. | CPCAT | |
|---------------|------|---------|-------|--|
| Poisson | 0.05 | 0.05 | 0.04 | |
| over under | 0.05 | 0.07 | 0.13 | |
| under | 0.06 | 0.05 | 0.03 | |

OECD proposed test: CPCAT III

- Properties of CPCAT
 - Falsely small p-values when data are overdispersed- i.e. in the most cases
 - Falsely large p-values when data are underdispersed- i.e. in some cases
 - Appropriate p-values when data are exactly Poisson distributed i.e. in rare case

OECD proposed test: CPCAT IV

• Empirical power (Without FWER control(); max power)

| True NOEC | Distrib. | MLT | Nonpar. | CPCAT |
|-----------|----------|------|---------|--------|
| 3 | Poisson | 0.76 | 0.82 | 0.82 |
| 3 | over | 0.58 | 0.60 | (0.69) |
| 3 | under | 0.83 | 0.88 | 0.67 |
| 2 | Poisson | 0.89 | 0.93 | 0.91 |
| 2 | over | 0.76 | 0.83 | (0.88) |
| 2 | under | 0.90 | 0.94 | 0.90 |
| 1 | Poisson | 0.92 | 0.95 | 0.96 |
| 1 | over | 0.86 | 0.94 | (0.96) |
| 1 | under | 0.94 | 0.95 | 0.97 |

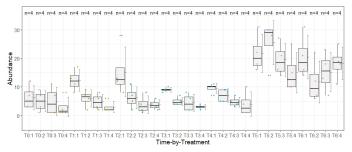
My advice: do not use CPCAT for routine analysis

Alternatives I

- Alternative I: Dunnett test, modified against variance heterogeneity using Welch-type df's [3]
- Alternative II: Nonparametric Dunnett-type test based on global ranks for relative effect sizes [5]. Can be used for both not-rare abundances and biomass
- Alternative III: Dunnett-type test based on most likely transformations sensitive for location/scale/shape effects [4]
- Use simultaneous two-sided $(1-2\alpha)$ confidence intervals: i) proof of hazard and safety, ii) decreasing effect at any monitoring time, possible followed by an increase later

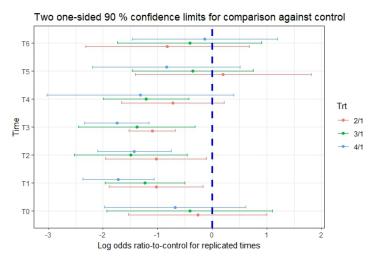
A joint approach: abundance example I

 The abundances of Ascidea in a complete randomized field trial using control and three concentrations (1,2,3,4)



- Nonparametric Dunnett-type test based on global ranks for relative effect sizes [5].
- The confidence intervals for log odds ratios versus control as effect size are presented for each sampling time (including pre-sampling T0):

A joint approach: abundance example II



A joint approach: abundance example III

- At pre-sampling time T0 near-to-equivalence can be concluded, i.e. no serious randomization bias occurred
- Already at T1 a significant decrease of the abundance in each of the concentration occurred, lasting until T3
- Starting at T4 a recovery effect can be observed, which becomes more pronounced at T5, 6.

Take home message I

- Estimate NOEC or BMD. The 1st is data- and design-dependent, the 2nd requires species-specific benchmark thresholds (BMR)
- Neither use CP-CAT nor Dunnett original test (nor any of the OECD proposed tests)
- Use the nonparametric Dunnett-type procedure: robust against variance heterogeneity and overdispersion
- Use confidence limits
- Further issue: KOVAR using pre-sampling data (under work)
- Related R-code available

Appendix: How to model the various sampling times T_t ? I

- First impuls: repeated measures analyses by mixed effect model or summarizing approaches (AUC) or even multivariate tests. NO!
- Second: using both T₀ AND T_t -separately
 - ► T₀ approach I): OECD recommends- demonstrate in each treatment group (before substance administration) a sufficient abundance. Using one-sided CI
 - ▶ T_0 approach II): OECD recommends- demonstrate no bias between treatment groups. Dunnett-type equivalence approach. Using 2-sided $(1-2\alpha)$ Cl's
 - Possible T₀ approach III): KOVAR
- Main objective: demonstrated a possible DECREASING abundance at any sampling time (will be species and ... dependent). Multiplicity-adjusted approaches are possible, but rather conservative (many t's, $n_i = 6$)- performed 1-sided CI unadjusted

Appendix: How to model the various sampling times T_t ? II

- Already such a claim for at least ONE species (and/or its aggregations) could be a final outcome of the trial. BUT
- Analyze for a possible recovery, i.e. an following increase of abundance for non-inferiority up to superiority- again by means of an one-sided CI
- Do both together: by two-sided $(1 2\alpha)$ Cl's
- See the example in a minute

References I

- [1] Joel E. Cohen. Sum of a random number of correlated random variables that depend on the number of summands. *Am. Stat.*, 73(1):56–60, 2019.
- [2] Benjamin Daniels. Application of the closure principle computational approach test to assess ecotoxicological field studies: Comparative analysis using earthworm field test abundance data. Environmental Toxicology and Chemistry, 40(6):1750–1760, 2021.
- [3] M. Hasler and L. A. Hothorn. Multiple contrast tests in the presence of heteroscedasticity. *Biometrical Journal*, 50(5):793–800, October 2008.
- [4] T. Hothorn, L. Most, and P. Buhlmann. Most likely transformations. Scandinavian Journal of Statistics, 45(1):110–134, March 2018.
- [5] F. Konietschke and L. A. Hothorn. Rank-based multiple test procedures and simult. confidence intervals. E. J. Stat., 6:738–759, 2012.